# CSE 548: Algorithms 

Randomized Algorithms
R. Sekar

## Example 1: Routing

- What is the best way to route a packet from $X$ to $Y$, esp. in high speed, high volume networks

A: Pick the shortest path from $X$ to $Y$
B: Send the packet to a random node $Z$, and let $Z$ route it to $Y$ (possibly using a shortest path from $Z$ to $Y$ )

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- Valiant showed in 1981 that surprisingly, B works better!
- Turing award recipient in 2010


## Example 2: Transmitting on shared network

- What is the best way for $n$ hosts to share a common a network?

A: Give each host a turn to transmit
B: Maintain a queue of hosts that have something to transmit, and use a FIFO algorithm to grant access
C: Let every one try to transmit. If there is contention, use random choice to resove it.

- Which choice is better?


## Topics

| 1. Intro | Caching |
| :--- | :---: |
| 2. Probability Basics | Hashing |
| Discrete Probability | Universal/Perfect hash |
| Coupon Collection | Closest pair |
| Birthday | 4. Probabilistic Algorithms |
| Balls and Bins | Bloom filter |
| 3. Taming distribution | Rabin-Karp |
| Quicksort | Prime testing |

## Simplify, Decentralize, Ensure Fairness

- Randomization can often:
- Enable the use of a simpler algorithm
- Cut down the amount of book-keeping
- Support decentralized decision-making
- Ensure fairness
- Examples:

Media access protocol: Avoids need for coordination - important here, because coordination needs connectivity!
Load balancing: Instead of maintaining centralized information about processor loads, dispatch jobs randomly.
Congestion avoidance: Similar to load balancing

## Set Theory and Probability

- A countable sample space $\mathcal{S}$ is a nonempty countable set.
- An outcome $\omega$ is an element of $\mathcal{S}$.
- A probability function $\operatorname{Pr}: \mathcal{S} \longrightarrow \mathbb{R}$ is a total function such that
- $\operatorname{Pr}[\omega] \geq 0$ for all $\omega \in \mathcal{S}$, and
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- $\operatorname{Pr}[\omega] \geq 0$ for all $\omega \in \mathcal{S}$, and
- $\sum_{\omega \in \mathcal{S}} \operatorname{Pr}[\omega]=1$
- An event $E$ is a subset of $\mathcal{S}$. Its probability is given by:

$$
\operatorname{Pr}[E]=\sum_{\omega \in E} \operatorname{Pr}[\omega]
$$

## Probability Rules from Set Theory

Many probability rules follow from the rules on set cardinality
Sum Rule: If $E_{0}, E_{1}, \ldots, E_{n}, \ldots$ are pairwise disjoint events, then

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Inclusion-Exclusion:

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\operatorname{Pr}[A \cup B]=\operatorname{Pr}[A]+\operatorname{Pr}[B]-\operatorname{Pr}[A \cap B]
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Union Bound: $\operatorname{Pr}[A \cup B] \leq \operatorname{Pr}[A]+\operatorname{Pr}[B]$

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Union Bound: $\operatorname{Pr}[A \cup B] \leq \operatorname{Pr}[A]+\operatorname{Pr}[B]$
Monotonicity: $A \subseteq B \rightarrow \operatorname{Pr}[A] \leq \operatorname{Pr}[B]$

## Uniform Probability Spaces

A finite probability space $\mathcal{S}$ said to be uniform if $\operatorname{Pr}[\omega]$ is the same for all $\omega$. In such spaces:

$$
\operatorname{Pr}[E]=\frac{|E|}{|\mathcal{S}|}
$$

We often this assumption.

## Conditional Probability

- Probability of an event under a condition
- The condition limits consideration to a subset of outcomes
- Consider this subset (rather than whole of $\mathcal{S}$ ) as the space of all possible outcomes

$$
\operatorname{Pr}[X \mid Y]=\frac{\operatorname{Pr}[X \cap Y]}{\operatorname{Pr}[Y]}
$$

## Extending Probability Rules for Conditional Probability

Product Rule 2: $\operatorname{Pr}\left[E_{1} \cap E_{2}\right]=\operatorname{Pr}\left[E_{1}\right] \cdot \operatorname{Pr}\left[E_{2} \mid E_{1}\right]$

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Total Probability Law: $\operatorname{Pr}[A]=\operatorname{Pr}[A \mid E] \cdot \operatorname{Pr}[E]+\operatorname{Pr}[A \mid \bar{E}] \cdot \operatorname{Pr}[\bar{E}]$
Total Probability Law 2: If $E_{i}$ are mutually disjoint and $\operatorname{Pr}\left[\bigcup E_{i}\right]=1$ then

$$
\operatorname{Pr}[A]=\sum \operatorname{Pr}\left[A \mid E_{i}\right] \cdot \operatorname{Pr}\left[E_{i}\right]
$$

Inclusion-Exclusion: $\operatorname{Pr}[A \cup B \mid C]=\operatorname{Pr}[A \mid C]+\operatorname{Pr}[B \mid C]-\operatorname{Pr}[A \cap B \mid C]$

## Independence

- An event $A$ is independent of $B$ iff the following (equivalent) conditions hold:
- $\operatorname{Pr}[A \mid B]=\operatorname{Pr}[A]$
- $\operatorname{Pr}[A \cap B]=\operatorname{Pr}[A] \cdot \operatorname{Pr}[B]$
- $B$ is independent of $A$
- Often, independence is an assumption.
- Definition can be generalized to 3 (or $n$ ) events. Events $E_{1}, E_{2}$ and $E_{3}$ a are mutually independent iff all of the following hold:
- $\operatorname{Pr}\left[E_{1} \cap E_{2}\right]=\operatorname{Pr}\left[E_{1}\right] \cdot \operatorname{Pr}\left[E_{2}\right]$
- $\operatorname{Pr}\left[E_{2} \cap E_{3}\right]=\operatorname{Pr}\left[E_{2}\right] \cdot \operatorname{Pr}\left[E_{3}\right]$
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## Coupon Collector Problem

- Suppose that your favorite cereal has a coupon inside. There are $n$ types of coupons, but only one of them in each box. How many boxes will you have to buy before you can expect to have all of the $n$ types?
- What is your guess?


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- What is your guess?
- Let us work out the expectation. Let us say that you have so far $j-1$ types of coupons, and are now looking to get to the $j$ th type. Let $X_{j}$ denote the number of boxes you need to purchase before you get the $j+1$ th type.


## Coupon Collector Problem

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## Coupon Collector Problem

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- We have all $n$ types when we finish the $X_{n-1}$ phase:

$$
E[X]=\sum_{i=0}^{n-1} E\left[X_{j}\right]=\sum_{i=0}^{n-1} n /(n-j)=n H(n)
$$

- Note $H(n)$ is the harmonic sum, and is bounded by $\ln n$


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- Note $H(n)$ is the harmonic sum, and is bounded by $\ln n$
- Perhaps unintuitively, you need to buy In $n$ cereal boxes to obtain one useful coupon.


## Birthday Paradox

- What is the smallest size group where there are at least two people with the same birthday?
- 365
- 183
- 61
- 25


## Birthday Problem

- The probability that two students have different birthdays: $\frac{364}{365}$
- In a class of $n$, there are $\binom{n}{2}$ pairs of students to consider.
- If we assume that whether one pair shares a birthday is independent of another, we can simply multiply these probabilities

$$
\operatorname{Pr}(\text { no two persons with same birthday }) \approx\left(\frac{364}{365}\right)^{\binom{n}{2}} \approx\left(\frac{364}{365}\right)^{n^{2} / 2}
$$

- For $n=44$, this formula yields a probability of $7 \%$
- $n=23$ is enough to have better than even chance of finding two with the same birthday.


## Birthday Problem: More Accurate Approach

- What is the probability of finding two people with the same birthday in this class?
- There are $365^{n}$ possible sequences of birthdays for $n$ people
- We assume these are all equally likely
- Number of sequences without repetition: $365 \cdot 364 \cdots(365-(n-1))$
- Probability that no two of $n$ persons have same birthday:

$$
\frac{365}{365} \cdot \frac{365-1}{365} \cdots \frac{365-(n-1)}{365}=\left(1-\frac{0}{365}\right)\left(1-\frac{1}{365}\right) \cdots\left(1-\frac{n-1}{365}\right)
$$

- Use the approximation $(1-x)<e^{-x}$ to derive an upper bound:
$\operatorname{Pr}$ (no two persons with same birthday) $<e^{0} \cdot e^{-\frac{1}{365}} \cdot e^{-\frac{n-1}{365}}=e^{\frac{-1}{365} \sum_{i=1}^{n-1} i}=e^{\frac{-n(n-1)}{2+365}}$
- For $n=44$, this evaluates to $7.5 \%$


## Birthday Paradox Vs Coupon Collection

- Two sides of the same problem

Coupon Collection: What is the minumum number of samples needed to cover every one of $N$ values
Birthday problem: What is the maximum number of samples that can avoid covering any value more than once?

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- So, if we want enough people to ensure that every day of the year is covered as a birthday, we will need $365 \ln 365 \approx 2153$ people!
- Almost 100 times as many as needed for one duplicate birthday!


## Balls and Bins

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- Birthday problem
- What should $m$ be to have at least one ball per bin?
- Coupon collection
- What is the maximum number of balls in any bin?
- Such problems arise in load-balancing, hashing, etc.


## Balls and Bins: Max Occupancy

- Probability $p_{1, k}$ that the first bin receives at least $k$ balls:
- Choose $k$ balls in $\binom{m}{k}$ ways
- These $k$ balls should fall into the first bin: prob. is $(1 / n)^{k}$
- Other balls may fall anywhere, i.e., probability $1:^{1}$

$$
\binom{m}{k}\left(\frac{1}{n}\right)^{k}=\frac{m \cdot(m-1) \cdots(m-k+1)}{k!n^{k}} \leq \frac{m^{k}}{k!n^{k}}
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- Let $m=n$, and use Sterling's approx. $k!\approx \sqrt{2 \pi k}(k / e)^{k}$ :

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P_{k}=\sum_{i=1}^{n} p_{i, k} \leq n \cdot \frac{1}{k!} \leq n \cdot\left(\frac{e}{k}\right)^{k}
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- Some arithmetic simplification will show that $P_{k}<1 / n$ when

$$
k=\frac{3 \ln n}{\ln \ln n}
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[^1]
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- Expected number of empty bins: $n e^{-m / n}$
- Max. balls in any bin when $m=n$ :

$$
\Theta(\ln n / \ln \ln n)
$$

- This is a probabilistic bound: chance of finding any bin with higher occupancy is $1 / n$ or less.
- Note that the absolute maximum is $n$.


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- thus, expected complexity of randomized quicksort is given by:

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T(n)=n+\frac{1}{n} \sum_{i=1}^{n-1}(T(i)+T(n-i))
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Summary: Input need not be random

- Expected $O(n \log n)$ performance comes from externally forced randomness in picking the pivot


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- But we can't predict the future!
- Result: many candidates for eviction. How can be avoid making bad (worst-case) choices repeatedly, even if input behaves badly?
- Approach: pick one of the candidates at random!


## Hash Tables

- A data structure for implementing:

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- Support expected $O(1)$ time lookup, insert, and delete
- Hash table entries may be:
fat: store a pair (key, object)
lean: store pointer to object containing key
- Two main questions:
- How to avoid $O(n)$ worst case behavior?
- How to ensure average case performance can be realized for arbitrary distribution of keys?


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- Sometimes, we fold the mod operation into $h$.
- Array elements typically called buckets
- Collisions bound to occur since $s \ll|\mathcal{U}|$
- Either $h(x)=h(y)$, or
- $h(x) \neq h(y)$ but $h(x) \equiv h(y)(\bmod s)$


## Collisions in Hash tables

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- If $\alpha=1$, what is the maximum number of collisions to expect?
- Both questions can be answered from balls-and-bins results: $1 / \sqrt{n}$, and $O(\ln n / \ln \ln n)$
- Real world keys are not random. Your hash table implementation needs to achieve its performance goals independent of this distribution.


## Chained Hash Table

- Each bucket is a linked list.
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- What is the average search time, as a function of $\alpha$ ?


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- Any key that hashes to a bucket is inserted into that bucket.
- What is the average search time, as a function of $\alpha$ ?
- It is $1+\alpha$ if:
- you assume that the distribution of lookups is independent of the table entries, OR,
- the chains are not too long (i.e., $\alpha$ is small)


## Open addressing

- If there is a collision, probe other empty slots

Linear probing: If $h(x)$ is occupied, try $h(x)+i$ for $i=1,2, \ldots$
Binary probing: Try $h(x) \oplus i$, where $\oplus$ stands for exor.
Quadratic probing: For ith probe, use $h(x)+c_{1} i+c_{2} i^{2}$

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- Criteria for secondary probes

Completeness: Should cycle through all possible slots in table Clustering: Probe sequences shouldn't coalesce to long chains Locality: Preserve locality; typically conflicts with clustering.

## Open addressing

- If there is a collision, probe other empty slots Linear probing: If $h(x)$ is occupied, try $h(x)+i$ for $i=1,2, \ldots$ Binary probing: Try $h(x) \oplus i$, where $\oplus$ stands for exor.
Quadratic probing: For $i$ th probe, use $h(x)+c_{1} i+c_{2} i^{2}$
- Criteria for secondary probes

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- Average search time can be $O\left(1 /(1-\alpha)^{2}\right)$ for linear probing, and $O(1 /(1-\alpha))$ for quadratic probing.


## Chaining Vs Open Addressing

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- Clustering causes more collisions w/ open addressing for same $\alpha$
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- Not all quadratic probes cover all slots (but some can)


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- Exercise: How to ensure amortized $O(1)$ cost when you resize up as well as down?


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Universal hashing
Perfect hashing

## Universal Hashing

- No single hash function can be good on all inputs
- Any function $\mathcal{U} \rightarrow\{1, \ldots, n\}$ must map $|\mathcal{U}| / n$ inputs to same value! Note: $|\mathcal{U}|$ can be much, much larger than $n$.


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## Definition

A family of hash functions $\mathcal{H}$ is universal if

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Meaning: If we pick $h$ at random from the family $\mathcal{H}$, then, probability of collisions is the same for any two elements.

Contrast with non-universal hash functions such as

$$
h(x)=a x \bmod n, \quad(a \text { is chosen at random })
$$

Note $y$ and $y+k n$ collide with a probability of 1 for every $a$.

## Universal Hashing Using Multiplication

Observation (Multiplication Modulo Prime)
If $p$ is a prime and $0<a<p$

- $\{1 a, 2 a, 3 a, \ldots,(p-1) a\}=\{1,2, \ldots, p-1\}(\bmod p)$
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## Prime multiplicative hashing

Let the key $x \in \mathcal{U}, p>|\mathcal{U}|$ be prime, and $0<r<p$ be random. Then

$$
h(x)=(r x \bmod p) \bmod n
$$

is universal.
Prove: $\operatorname{Pr}[h(x)=h(y)]=\frac{1}{n}, \quad$ for $x \neq y$

## Universality of prime multiplicative hashing

- Need to show $\operatorname{Pr}[h(x)=h(y)]=\frac{1}{n}$, for $x \neq y$
- $h(x)=h(y)$ means $(r x \bmod p) \bmod n=(r y \bmod p) \bmod n$
- Note $a \bmod n=b \bmod n$ means $a=b+k n$ for some integer $k$. Using this, we eliminate $\bmod n$ from above equation to get:

$$
\begin{aligned}
r x \bmod p & =k n+r y \bmod p, \text { where } k \leq\lfloor p / n\rfloor \\
r x & \equiv k n+r y(\bmod p) \\
r(x-y) & \equiv k n(\bmod p) \\
r & \equiv k n(x-y)^{-1}(\bmod p)
\end{aligned}
$$

- So, $x, y$ collide if $r=n(x-y)^{-1}, 2 n(x-y)^{-1}, \ldots,\lfloor p / n\rfloor n(x-y)^{-1}$
- In other words, $x$ and $y$ collide for $p / n$ out of $p$ possible values of $r$, i.e., collision probability is $1 / n$


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- When $|\mathcal{U}|<2^{w}, n=2^{l}$ and $a$ an odd random number

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(a*x) >> (WORDSIZE-HASHBITS)


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- Can be implemented efficiently if $w$ is the wordsize:

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\left(a^{*} \mathrm{x}\right) ~ \gg \text { (WORDSIZE-HASHBITS) }
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- Scheme is near-universal: collision probability is $O(1) / 2^{l}$


## Prime Multiplicative Hash for Vectors

Let $p$ be a prime number, and the key $x$ be a vector $\left[x_{1}, \ldots, x_{k}\right]$ where $0 \leq x_{i}<p$. Let

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h(x)=\sum_{i=1}^{k} r_{i} x_{i}(\bmod p)
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If $0<r_{i}<p$ are chosen at random, then $h$ is universal.

- Strings can also be handled like vectors, or alternatively, as a polynomial evaluated at a random point $a$, with $p$ a prime:

$$
h(x)=\sum_{i=0}^{l} x_{i} a^{i} \bmod p
$$

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- Since $y_{i}, x_{i}<p$, it is easy to see from this equation that the collision-causing value of $r_{i}$ is distinct for distinct $y_{i}$.
- Viewed another way, exactly one of $p$ choices of $r_{i}$ would cause a collision between $x_{i}$ and $y_{i}$, i.e., $\operatorname{Pr}_{h}[h(x)=h(y)]=1 / p$


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Static: Pick a hash function (or set of functions) that avoids collisions for a given set of keys

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Don't forget: more memory usually means less performance due to cache effects.
Approach 2: Use a secondary hash table for each bucket of size $n_{i}^{2}$, where $n_{i}$ is the number of elements in the bucket.
Uses only $O(n)$ storage, if $h$ is universal

## Hashing Summary

- Excellent average case performance
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- Excellent average case performance
- Pointer chasing is expensive on modern hardware, so improvement from $O(\log n)$ of binary trees to expected $O(1)$ for hash tables is significant.
- But all benefits will be reversed if collisions occur too often
- Universal hashing is a way to ensure expected average case even when input is not random.
- Perfect hashing can provide efficient performance even in the worst case, but the benefits are likely small in practice.


## Finding closest pair of points

Problem: Given a set of $n$ points in a $d$-dimensional space, identify the two that have the smallest Euclidean distance between them.


Applications: A central problem in graphics, vision, air-traffic control, navigation, molecular modeling, and so on.

[^2]
## Randomized Closest Pair: Key Ideas

- Divide the plane into small squares, hash points into them
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- Process the points in some random order
- Maintain min. distance $\delta$ among points processed so far.
- Update $\delta$ as more points are processed
- At any point, the "small squares" have a size of $\delta / 2$
- At most one point per square (or else points are closer than $\delta$ )
- Points closer than $\delta$ will at most be two squares from each other
- Only constant number of points to consider
- Requires rehashing all processed points when $\delta$ is updated.


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- Correctness is relatively clear, so we focus on performance
- Two main concerns


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- Total runtime can all be "charged" to insert operations,
- incl. those performed during rehashing
so we will focus on estimating inserts.


## Randomized Closest Pair: \# of Inserts

## Theorem

If random variable $X_{i}$ denotes the likelihood of needing to rehash after processing $k$ points, then

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X_{i} \leq \frac{2}{i}
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- Since points are processed in random order, there is a $2 / i$ probability that $p_{i}$ is one of $p$ or $q$


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Look Ma! I have a linear-time randomized closest pair algorithm-And it is not even probabilistic!

## Probabilistic Algorithms

- Algorithms that produce the correct answer with some probability
- By re-running the algorithm many times, we can increase the probability to be arbitrarily close to 1.0 .


## Bloom Filters

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- What if you want to store very large keys?
- Radical idea: Don't store the key in the table!
- Potentially w-fold space reduction


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Images from Wikipedia Commons

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- No deletions possible in the current algorithm.


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- Complementing, the prob. $p$ that a certain bit is set is $1-e^{-k r / m}$


## Bloom Filters: False positives

- Prob. that a bit is not set by $h_{1}$ on inserting a key is $(1-1 / m)$
- The probability it is not set by any $h_{i}$ is $(1-1 / m)^{k}$
- The probability it is not set after $r$ key inserts is $(1-1 / m)^{k r} \approx e^{-k r / m}$
- Complementing, the prob. $p$ that a certain bit is set is $1-e^{-k r / m}$
- For a false positive on a key $y$, all the bits that it hashes to should be a 1 . This happens with probability

$$
\left(1-e^{-k r / m}\right)^{k}=(1-p)^{k}
$$

## Bloom Filters

- Note: $n=m / r$ is the storage (in bits) used per key.
- So, we can rewrite the FP equation as:

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- A Bloom filter that uses just 8 bits per key to store an arbitrary sized key will have an FP rate of $2 \%$
- Important: Bloom filters can be used as a prefilter, e.g., if actual keys are in secondary storage (e.g., files or internet repositories)


## Using arithmetic for substring matching

Problem: Given strings $T$ [1..n] and $P[1 . . m]$, find occurrences of $P$ in $T$ in $O(n+m)$ time.

Idea: To simplify presentation, assume $P, T$ range over [0-9]

- Interpret $P[1 . . m]$ as digits of a number

$$
p=10^{m-1} P[1]+10^{m-2} P[2]+\cdots 10^{m-m} P[m]
$$

- Similarly, interpret $T[i . .(i+m-1)]$ as the number $t_{i}$
- Note: $P$ is a substring of $T$ at $i$ iff $p=t_{i}$
- To get $t_{i+1}$, shift $T[i]$ out of $t_{i}$, and shift in $T[i+m]$ :

$$
t_{i+1}=\left(t_{i}-10^{m-1} T[i]\right) \cdot 10+T[i+m]
$$

We have an $O(n+m)$ algorithm. Almost: we still need to figure out how to operate on $m$-digit numbers in constant time!

## Rabin-Karp Fingerprinting

## Key Idea

- Instead of working with m-digit numbers,
- perform all arithmetic modulo a random prime number $q$,
- where $q>m^{2}$ fits within wordsize
- All observations made on previous slide still hold
- Except that $p=t_{i}$ does not guarantee a match
- Typically, we expect matches to be infrequent, so we can use $O(m)$ exact-matching algorithm to confirm probable matches.


## Carter-Wegman-Rabin-Karp Algorithm

Difficulty with Rabin-Karp: Need to generate random primes, which is not an efficient task.
New Idea: Make the radix random, as opposed to the modulus

- We still compute modulo a prime $q$, but it is not random.

Alternative interpretation: We treat $P$ as a polynomial

$$
p(x)=\sum_{i=1}^{m} P[m-i] \cdot x^{i}
$$

and evaluate this polynomial at a randomly chosen value of $x$
Like any probabilistic algorithm we can increase correctness probability by repeating the algorithm with different randoms.

- Different prime numbers for Rabin-Karp
- Different values of $x$ for CWRK


## Carter-Wegman-Rabin-Karp Algorithm

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Random choice does not imply high probability of being right.

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- Arithmetic modulo prime defines a field, so an $m$ th degree polynomial has $m+1$ roots.
- Thus, $(m+1) / q$ of the $q$ (recall $q$ is the prime number used for performing modulo arithmetic) possible choices of $x$ will result in a false match, i.e., probability of false positive $=(m+1) / q$


## Primality Testing

## Fermat's Theorem <br> $a^{p-1} \equiv 1(\bmod p)$

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- Canceling out $(p-1)$ ! from both sides, we have the theorem!


## Primality Testing

- Given a number $N$, we can use Fermat's theorem as a probabilistic test to see if it is prime:
- if $a^{N-1} \not \equiv 1(\bmod N)$ then $N$ is not prime
- Repeat with different values of $a$ to gain more confidence


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- For other numbers, we can show that the above procedure works with probability 0.5


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- Thus, for every $b$ for which Fermat's test is satisfied, there exists a $c$ that does not satisfy it.
- Moreover, since $a$ is relatively prime to $N, a b \not \equiv a b^{\prime}$ unless $b \equiv b^{\prime}$.


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- Thus, for every $b$ for which Fermat's test is satisfied, there exists a $c$ that does not satisfy it.
- Moreover, since $a$ is relatively prime to $N, a b \not \equiv a b^{\prime}$ unless $b \equiv b^{\prime}$.
- Thus, at least half of the numbers $x<N$ relatively prime to $N$ will fail the test.


## Primality Testing



- When Fermat's test returns "prime" $\operatorname{Pr}[\mathrm{N}$ is not prime $]<0.5$
- If Fermat's test is repeated for $k$ choices of $a$, and returns "prime" in each case, $\operatorname{Pr}[N$ is not prime $]<0.5^{k}$
- In fact, 0.5 is an upper bound. Empirically, the probability has been much smaller.


## Primality Testing



- Empirically, on numbers less than 25 billion, the probability of Fermat's test failing to detect non-primes (with $a=2$ ) is more like 0.00002
- This probability decreases even more for larger numbers.


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## Lagrange's Prime Number Theorem

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- $O\left(\log ^{2} N\right)$ multiplications on $\log N$ bit numbers
- If $N$ is not prime, should we try $N+1, N+2, \ldots$ instead of generating a new random number?


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- If $N$ is not prime, should we try $N+1, N+2, \ldots$ instead of generating a new random number?
- No, it is not easy to decide when to give up.


## Rabin-Miller Test

- Works on Carmichael's numbers
- For prime number test, we consider only odd $N$, so $N-1=2^{t} u$ for some odd $u$
- Compute

$$
a^{u}, a^{2 u}, a^{4 u}, \ldots, a^{2^{t} u}=a^{N-1}
$$

- If $a^{N-1}$ is not 1 then we know $N$ is composite.
- Otherwise, we do a follow-up test on $a^{u}, a^{2 u}$ etc.
- Let $a^{2^{r} u}$ be the first term that is equivalent to 1 .
- If $r>0$ and $a^{2^{r-1} u} \not \equiv-1$ then $N$ is composite
- This combined test detects non-primes with a probability of at least 0.75 for all numbers.


[^0]:    ${ }^{1}$ This is actually an upper bound, as there can be some double counting.

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[^2]:    Images from Wikipedia Commons

